- Joint typical sequences
- Covering and Packing Lemmas
- Channel Coding Theorem
- Capacity of Gaussian channel
- Capacity of additive white Gaussian channel
- Forward proof of Channel Coding Theorem

This time

- Converse Proof of Channel Coding Theorem
- Non-white Gaussian Channel
- Rate-distortion problems
- Rate-distortion Theorem

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To continue the converse proof, we will need to introduce a simple result from Fano

Fano's inequality

Denote $Pr(error) = P_e = Pr(M \neq \hat{M})$, then $H(M|Y^N) \leq 1 + P_eH(M)$ Intuitively, if $P_e \rightarrow 0$, on average we will know M for certain given y and thus $\frac{1}{N}H(M|Y^N) \rightarrow 0$

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Proof: Let $E = I(M \neq \hat{M})$, then

 $H(M|Y^N) = H(M, E|Y^N) - H(E|Y^N, M)$

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$$H(M|Y^N) = H(M, E|Y^N) - H(E|Y^N, M)$$

= $H(M, E|Y^N) = H(E|Y^N) + H(M|Y^N, E)$

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$$\begin{aligned} H(M|Y^{N}) &= H(M, E|Y^{N}) - H(E|Y^{N}, M) \\ &= H(M, E|Y^{N}) = H(E|Y^{N}) + H(M|Y^{N}, E) \\ &\leq H(E) + H(M|Y^{N}, E) \\ &\leq 1 + P(E=0)H(M|Y^{N}, E=0) + P(E=1)H(M|Y^{N}, E=1) \end{aligned}$$

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$$R = \frac{H(M)}{N} = \frac{1}{N} \left[I(M; Y^N) + H(M|Y^N) \right]$$

Image: A matrix and a matrix

$$R = \frac{H(M)}{N} = \frac{1}{N} \left[I(M; Y^N) + H(M|Y^N) \right] \leq \frac{1}{N} \left[I(X^N; Y^N) + H(M|Y^N) \right]$$

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- But sometimes noise power can be different for different band, consequently, "color" channels
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- Without loss of generality, let's consider the discrete approximation, parallel Gaussian channel

Consider that we have K parallel channels (K bands) and the corresponding noise powers are σ²₁, σ²₂, · · · , σ²_K

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- Therefore, for the k-th channel, we can transmit $\frac{1}{2}\log\left(1+\frac{P_k}{\sigma_k^2}\right)$ bits per channel use
- So our goal is to assign $P_1, P_2, \cdots, P_K \ge 0$ $(\sum_{k=1}^K P_k \le P)$ such that the total capacity

$$\sum_{k=1}^{K} \frac{1}{2} \log \left(1 + \frac{P_k}{\sigma_k^2}\right)$$

is maximize

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$$\mu \left(\sum_{k=1}^{K} P_k - P \right) = 0, \qquad \lambda_k P_k = 0, \forall k$$

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Since $\lambda_i P_i = 0$, for $P_i > 0$, we have $\lambda_i = 0$ and thus

$$P_i + \sigma_i^2 = \frac{1}{2\mu}$$
Capacity of parallel channels

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Lecture 12 Capacity of non-white Gaussian channels

Water-filling interpretation



From $P_i + \sigma_i^2 = const$, power can be allocated intuitively as filling water to a pond (hence "water-filling")

Example



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- There is an apparent rate (bits per sample) and distortion (fidelity) trade-off. We expect that needed rate is smaller if we allow a lower fidelity (higher distortion). What we are really interested in is a rate-distortion function

$$m \in \{1, 2, \cdots, M\}$$

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- Note that we have a freedom to pick p(x̂|x) such that E[d(X̂^N, X^N)] (less than or) equal to the desired D

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- How is it related to the distortion though?
- Note that we have a freedom to pick $p(\hat{x}|x)$ such that $E[d(\hat{X}^N, X^N)]$ (less than or) equal to the desired \mathcal{D}
- Therefore given \mathcal{D} , the rate-distortion function is simply

$$R(\mathcal{D}) = \min_{p(\hat{x}|x)} I(\hat{X}; X)$$

such that $E[d(\hat{X}^N, X^N)] \leq \mathcal{D}$

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- We need to introduce a distortion measure first. Note that we have two types of errors: taking head as tail and taking tail as head. A natural measure will just weights both error equally

$$d(X = H, \hat{X} = T) = d(X = T, \hat{X} = H) = 1$$

$$d(X = H, \hat{X} = H) = d(X = T, \hat{X} = T) = 0$$

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$$d(X = H, \hat{X} = T) = d(X = T, \hat{X} = H) = 1$$

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• If rate is > 1 bit, we know that distortion is 0. How about rate is 0, what distortion suppose to be?

- Let's try to compress outcome from a fair coin toss
- We know that we need 1 bit to compress the outcome losslessly, what if we have only 0.5 bit per sample?
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- If decoders know nothing, the best bet will be just always decode head (or tail). Then D = E[d(X, H)] = 0.5

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For 0 < D < 0.5, denote Z as the prediction error such that $X = \hat{X} + Z$. Note that

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0.1

$$d(\hat{X},X) = (\hat{X} - X)^2$$

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Given distortion constraint \mathcal{D} , we can find scheme such that the require rate is no bigger than

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where the \hat{X} introduced by $p(\hat{x}|x)$ should satisfy $E[d(X, \hat{X})] \leq D$

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Note that the code rate is $\frac{\log 2^{NR}}{N} = R$ as desired

We say joint typical sequences x^N and \hat{x}^N are distortion typical $((x^N, \hat{x}^N) \in \mathcal{A}_{d,\epsilon}^N)$ if $|d(x^N, \hat{x}^N) - E[d(X, \hat{X})]| \le \epsilon$

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- Consequently, $(1 \delta)2^{N(H(X,\hat{X}) \epsilon)} \le |\mathcal{A}_{d,\epsilon}^N| \le 2^{N(H(X,\hat{X}) + \epsilon)}$ as before
- For two independently drawn sequences \hat{X}^N and X^N , the probability for them to be distortion typical will be just the same as before. In particular, $(1 \delta)2^{-N(I(X;\hat{X}) 3\epsilon)} \leq Pr((X^N, \hat{X}^N) \in \mathcal{A}_{d,\epsilon}^N(X, \hat{X}))$

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Given input X^N , find out of the codewords the one that is jointly typical with X^N . And say, if the codeword is C(i), output index *i* to the decoder

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• First of all, the only point of failure lies on encoding, that is when the encoder cannot find a codeword jointly typical with X^N

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- By covering Lemma, encoding failure is neglible as long as $R > I(X; \hat{X})$
- If encoding is successful, C(i) and X^N should be distortion typical. Therefore, E[d(C(i); X^N)] ~ E[d(X̂, X)] ≤ D as desired

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In the proof, we need to use the convex property of $R(\mathcal{D})$. That is,

$$R(a\mathcal{D}_1 + (1-a)\mathcal{D}_2) \ge aR(\mathcal{D}_1) + (1-a)R(\mathcal{D}_2)$$

So we will digress a little bit to show this convex property first

Log-sum inequality

For any $a_1, \dots, a_n \ge 0$ and $b_1, \dots, b_n \ge 0$, we have $\sum_i a_i \log_2 \frac{a_i}{b_i} \ge \sum_i a_i \log_2 \frac{\sum_i a_i}{\sum_i b_i}.$

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Proof

We can define two distributions p(x) and q(x) with $p(x_i) = \frac{a_i}{\sum_i a_i}$ and $q(x_i) = \frac{b_i}{\sum_i b_i}$. Since p(x) and q(x) are both non-negative and sum up to 1, they are indeed valid probability mass functions.

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We can define two distributions p(x) and q(x) with $p(x_i) = \frac{a_i}{\sum_i a_i}$ and $q(x_i) = \frac{b_i}{\sum_i b_i}$. Since p(x) and q(x) are both non-negative and sum up to 1, they are indeed valid probability mass functions. Then, we have

Log-sum inequality

For any $a_1, \dots, a_n \ge 0$ and $b_1, \dots, b_n \ge 0$, we have $\sum_i a_i \log_2 \frac{a_i}{b_i} \ge \sum_i a_i \log_2 \frac{\sum_i a_i}{\sum_i b_i}.$

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For any four distributions $p_1(\cdot)$, $p_2(\cdot)$, $q_1(\cdot)$, and $q_2(\cdot)$, we have

 $\lambda_1 \mathsf{KL}(p_1 \| q_1) + \lambda_2 \mathsf{KL}(p_2 \| q_2) \geq \mathsf{KL}(\lambda_1 p_1 + \lambda_2 p_2 \| \lambda_1 q_1 + \lambda_2 q_2),$

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= $\lambda_1 \sum_{x \in \mathcal{X}} p_1(x) \log \frac{p_1(x)}{q_1(x)} + \lambda_2 \sum_{x \in \mathcal{X}} p_2(x) \log \frac{p_2(x)}{q_2(x)}$

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Convexity of I(X; Y) with respect to p(y|x)

For any random variables X and Y, I(X; Y) is a convex function of p(y|x) for a fixed p(x)

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I(X; Y) is concave with respect to p(x) for fixed p(y|x) though. A proof is given in Cover and Thomas and will be omitted here

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Proof

Let us write

$$(X; Y) = KL(p(x, y) || p(x) p(y))$$

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We want to show

 $\lambda f(p_1(y|x)) + (1-\lambda)f(p_2(y|x)) \geq f(\lambda p_1(y|x) + (1-\lambda)p_2(y|x))$

Continue from previous slide, we have

$$\lambda f(p_1(y|x)) + (1 - \lambda) f(p_2(y|x))$$

= $\lambda K L \Big(p(x) p_1(y|x) \Big\| p(x) \sum_x p(x) p_1(y|x) \Big)$
+ $(1 - \lambda) K L \Big(p(x) p_2(y|x) \Big\| p(x) \sum_x p(x) p_2(y|x) \Big)$

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Continue from previous slide, we have

$$\begin{split} \lambda f(p_{1}(y|x)) &+ (1-\lambda) f(p_{2}(y|x)) \\ &= \lambda K L \Big(p(x) p_{1}(y|x) \Big\| p(x) \sum_{x} p(x) p_{1}(y|x) \Big) \\ &+ (1-\lambda) K L \Big(p(x) p_{2}(y|x) \Big\| p(x) \sum_{x} p(x) p_{2}(y|x) \Big) \\ &\geq K L \Big(\lambda p(x) p_{1}(y|x) + (1-\lambda) p(x) p_{2}(y|x) \Big\| \lambda p(x) \sum_{x} p(x) p_{1}(y|x) \\ &+ (1-\lambda) p(x) \sum_{x} p(x) p_{2}(y|x) \Big) \end{split}$$

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Recall that $R(\mathcal{D}) = \min_{p(\hat{x}|x)} I(\hat{X}; X)$ with $E[d(X, \hat{X})] \leq \mathcal{D}$ We want to show that

 $R(\lambda \mathcal{D}_1 + (1-\lambda)\mathcal{D}_2) \leq \lambda R(\mathcal{D}_1) + (1-\lambda)R(\mathcal{D}_2)$

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Recall that $R(D) = \min_{p(\hat{x}|x)} I(\hat{X}; X)$ with $E[d(X, \hat{X})] \leq D$ We want to show that

$$\mathsf{R}(\lambda\mathcal{D}_1+(1-\lambda)\mathcal{D}_2)\leq\lambda\mathsf{R}(\mathcal{D}_1)+(1-\lambda)\mathsf{R}(\mathcal{D}_2)$$

Proof

Let $p_1^*(\hat{x}|x)$ and $p_2^*(\hat{x}|x)$ be the distributions that optimize $R(\mathcal{D}_1)$ and $R(\mathcal{D}_2)$. Let's try to time share between the two distributions.

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Recall that $R(D) = \min_{p(\hat{x}|x)} I(\hat{X}; X)$ with $E[d(X, \hat{X})] \leq D$ We want to show that

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with λ fraction of time with $(1-\lambda)$ fraction of time

where $\tilde{X} = \begin{cases} \hat{X}_1 \\ \hat{X}_2 \end{cases}$

$$p(x) \xrightarrow{X^{N}} Encoder \xrightarrow{m} Decoder \xrightarrow{} \hat{X}^{N}$$

 $NR \geq H(M)$

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$$\begin{array}{c|c} p(x) & X^N & Encoder & m \\ \hline & Decoder & \hat{X}^N \end{array}$$

 $NR \ge H(M) \ge H(M) - H(M|X^N) = I(M;X^N)$

$$p(x) \xrightarrow{X^{N}} Encoder \xrightarrow{m} Decoder \xrightarrow{} \hat{X}^{N}$$

 $NR \ge H(M) \ge H(M) - H(M|X^N) = I(M;X^N) \ge I(\hat{X}^N;X^N)$

 $= H(X^N) - H(X^N | \hat{X}^N)$

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$$p(x) \xrightarrow{X^{N}} Encoder \xrightarrow{m} Decoder \xrightarrow{X^{N}} \hat{X}^{N}$$

$$NR \ge H(M) \ge H(M) - H(M|X^{N}) = I(M; X^{N}) \ge I(\hat{X}^{N}; X^{N})$$

$$= H(X^{N}) - H(X^{N}|\hat{X}^{N}) = \sum_{i=1}^{N} H(X_{i}) - \sum_{i=1}^{N} H(X_{i}|\hat{X}^{N}, X^{i-1})$$

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Image: A matrix and a matrix

$$p(x) \xrightarrow{X^{N}} Encoder \xrightarrow{m} Decoder \xrightarrow{} \hat{X}^{N}$$

$$NR \ge H(M) \ge H(M) - H(M|X^{N}) = I(M; X^{N}) \ge I(\hat{X}^{N}; X^{N})$$

$$= H(X^{N}) - H(X^{N}|\hat{X}^{N}) = \sum_{i=1}^{N} H(X_{i}) - \sum_{i=1}^{N} H(X_{i}|\hat{X}^{N}, X^{i-1})$$

$$\ge \sum_{i=1}^{N} H(X_{i}) - \sum_{i=1}^{N} H(X_{i}|\hat{X}_{i}) = \sum_{i=1}^{N} I(X_{i}; \hat{X}_{i})$$

-

Image: A matrix and a matrix

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$$\ge \sum_{i=1}^{N} R(E[d(X_{i}, \hat{X}_{i})]) = N\left(\frac{1}{N}\sum_{i=1}^{N} R(E[d(X_{i}; \hat{X}_{i})])\right)$$

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Image: A mathematical states of the state

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$$\ge NR\left(\frac{1}{N}\sum_{i=1}^{N} E[d(X_{i};\hat{X}_{i})]\right) = NR\left(E\left[\frac{1}{N}\sum_{i=1}^{N} d(X_{i};\hat{X}_{i})\right]\right)$$

Image: A matrix and a matrix

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$$= NR(E[d(X^{N}; \hat{X}^{N})]) \ge NR(D)$$

S. Cheng (OU-Tulsa)